

Profitability and cost shifting in government procurement contracts^{*}

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This draft: August 2014

Abstract: We examine whether contractors with cost plus contracts earn a higher profit, and whether the higher profit is associated with cost shifting behavior, using a unique set of data of federal procurement contracts between the years 2005 and 2010. Prior research often examines government contractor profitability without differentiating the contract types, which is a key determinant to the contractor's ability to cost-shift. We identify firms that are awarded with cost plus contracts in some years and without in other years. We find their profitability significantly increases during the years that cost plus contracts are awarded. We also find that these firms exhibit greater discretionary expenditures during the years with cost plus contracts, relative to the years without, which is consistent with cost shifting behavior. However, effective monitoring through the Cost Accounting Standards helps mitigate such behavior.

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1. Introduction

The U.S. government spends an enormous amount of money on public procurement. For example, the U.S. federal government reported \$391.3 billion and \$540.1 billion on procurement in the year 2005 and 2010, respectively.¹ Partially due to its significance in amount, popular opinions toward government procurement have been controversial. It is often believed that government contractors earn too high a profit at the tax payers' expense, and the government does not provide enough monitoring to discipline the contractors. To respond to the growing dissatisfaction against federal contractors, President Obama has proposed and implemented a series of new initiatives to further tighten the requirements for federal contractors. One of these new measures is to move away from cost plus contracts while adopting more fixed price contracts (Michaels and Cole, 2009).

Fixed price and cost plus are the two most common types of procurement contracts offered by the U.S. federal government to procurement contractors.² For a fixed price contract, a contractor provides a product or service to the government at a fixed price that is previously determined through the negotiation or bidding process. For a cost plus contract, the contract price is equal to a contractor's cost to produce the product or service plus a profit margin. That is, with a cost plus contract, the government reimburses the contractor for its *declared costs* to fulfill the contract, and pays it an additional fee or profit rate according to the negotiation. Therefore, if offered a cost plus contract, a contractor has an incentive to inflate its reported costs to earn a higher profit. Specifically, the contractor could use accounting discretions to shift costs from

¹ See: <http://www.usaspending.gov/>

² In practice, there can be hybrids of both types of contracts. For example, Rogerson (1992) describes four common types of contracts employed by the U.S. Department of Defense: (1) pure fixed price (2) pure cost reimbursement (3) incentive fixed price (i.e. realized costs are reimbursed up to a pre-specified threshold level) (4) incentive cost reimbursement (i.e. incentive fixed price contracts revert to pure cost reimbursement at a pre-specified threshold cost level). For our analysis, we refer to any contracts with revenue that is sensitive to the seller's cost as cost plus. In sensitivity analysis we ensure our results are robust to this empirical choice.

other segments of its business to the government business. Cost plus contracts are therefore often considered as “abuse-prone.”

In this paper, we use a unique set of data to examine whether contractors with cost plus contracts earn a higher profit, and whether the higher profit is associated with cost shifting behavior. To control for other confounding exogenous factors, we identify firms that are awarded with cost plus contracts in some years and without in other years. We find their profitability significantly increases during the years that cost plus contracts are awarded. We also find that these firms exhibit greater discretionary expenditures during the years with cost plus contracts, relative to the years without, consistent with cost shifting behavior. However, effective monitoring through the Cost Accounting Standards seems to help mitigate such behavior.

A vast theoretical literature explores the issues of optimal contracts and information problems in procurement setting (Laffont and Tirole 1986; McAfee and McMillan 1986; Rogerson 1992; Rogerson 1994; etc.). Since the government cannot costlessly observe the true cost incurred by a contractor, the contractor with a cost plus contract can manipulate the reported cost through cost inflation or cost shifting. Specifically, Rogerson (1992) demonstrates that a government contractor could over-allocate overhead cost into government contracts, while under-allocate overhead cost into commercial contracts, to strategically shift cost from commercial business to government business. Further, with the presence of such informational asymmetry, a first best solution can never be achieved. In contrast, fixed price contracts do not lead to such distortion of incentives as the contractor fully captures any profit/loss earned under the fixed price, regardless of its real cost or reported cost.

Empirical evidence on government contractors’ cost-shifting behavior has been scant, mainly due to the lack of data. Until 2004, information about which government contractors

received a cost plus contract was not publicly disclosed. A few empirical papers examine cost shifting and profitability among government contractors, but without differentiating the types of contracts. Thomas and Tung (1992) find evidence that department of defense contractors shift pension costs to the government. Lichtenberg (1992) finds government contractors earned excess profits on government contracts in the 1980s and suggests cost shifting as the reason. However, McGown and Vondrzyk (2002) hypothesize that a government contractor can only shift costs in a business segment with both government and commercial contracts. They fail to find significant differences in profitability between such business segments and their counterparts with only government or only commercial contracts. Therefore, they conclude that government contractors do not engage in cost shifting.

Prior research has also examined cost shifting in other settings. For example, Eldenburg and Soderstrom (1996) find that hospitals are engaged in cost shifting among payors using data from hospitals in the state of Washington. The cost shifting practice they focus on involves the hospital managers purposefully biasing budgeted information. Eldenburg and Kallapur (1997) examine the hospitals' response to a Medicare policy change in 1983. They find that the hospitals maximize their revenues by changing their patient mix and overhead. Eldenburg and Krishnan (2006) provide a comprehensive review on cost and incentive issues in the healthcare industry.

In this paper, we examine whether government contractors with cost plus contracts earn an unusually high profit; and if so, whether the high profitability is associated with cost shifting facilitated by accounting discretions. To address our research questions, we construct a database of 4.9 million federal procurement contracts (5,745 firm-years; 1,137 firms) between 2005 and

2010 from the Federal Procurement Data System website.³ For our main test, we identify U.S. publicly-traded government contractors that have been awarded with a cost plus contract in at least one year, and with only fixed price contracts in at least another one year during our sample period. The final sample consists of 258 firms, or 1,505 firm-years. This sample enables us to use the firm as its own control and mitigates concerns that firm characteristics drives differences in cost shifting and profitability. We find that our sample firms have greater discretionary expenditures as well as higher profitability during the years with cost plus contracts, relative to other years without cost plus contracts. This is consistent with the conventional wisdom that government contractors that are awarded with cost plus contracts engage in cost shifting to obtain higher profits.

In addition, we examine the effect of government monitoring through two cross-sectional tests. We identify contracts that are required to be compliant with the Cost Accounting Standards (CAS). When contracts (both fixed and cost plus) are subject to CAS coverage, the firm is required to provide a detailed disclosure about the accounting policies used to determine CAS related costs, thus facing increased monitoring efforts by the government to prevent cost shifting. We predict the association between cost plus contracts and profitability to decrease as the percentage of revenue subject to CAS increases. Our results are consistent with our predictions.

We conduct several sensitivity tests to rule out alternative explanations. First, we employ different samples to ensure our results are not sensitive to sample selection. For our main analysis, the firm serves as its own control, by using 1,505 firm-years (258 firms) with at least one year with a cost contract and one year without a cost contract over the sample period. Our results are robust to two alternative subsamples: (1) all contractors (5,745 firm-years, 1,137

³ https://www.fpds.gov/fpdsng_cms/index.php/en/

firms) and (2) all contractors with at least one cost contract (2,201 firm-years, 383 firms).

Second, for our main analyses, we refer to any contracts with revenue that is sensitive to the seller's cost as cost plus. However, in practice contracts are often hybrids that fall in the spectrum between cost plus and fixed price. Our results are robust to using a sample which includes only pure fixed price and pure cost reimbursement.

Please note that our findings do not endorse the view that cost plus contracts should be reduced or even eliminated in government procurement. Different forms of contracts serve different purposes and different contracting environments. Prior research (citexxx) on optimal contract form demonstrates that factors such as project-specific risk, market competition, contractor characteristics, and trust, are critical in determining the contract form. For example, when the nature of a project is long-term and highly uncertain, the government is more likely to adopt cost plus contract to share risk. Same is true is when the contractor firm has low risk tolerance because it is young in age, small in size, and/or highly-leveraged. Fixed price contracts often imply higher risk for the contractor, thus applicable only when the project outcome is certain or when the contractor has high risk tolerance. Further, the informational rent earned by government contractors is inevitable when there is informational asymmetry. The government can never obtain the contractors' true cost information because frequent audits are costly. The government must trade off the potential gains recovered from these audits and the monitoring cost.

Our paper makes several contributions to the literature. First, we are the first paper that examines government procurement using contract-level data, which provides a cleaner setting than prior studies. McGowan and Vondra (2002) do not differentiate contract types and fail to find evidence of higher profitability for defense contractors. This result could be due to the fact

that, for our sample, only 4% of contracts are cost plus and only 37% of government contractors have cost plus contracts. Not differentiating the contract types could bias against finding differential profitability because only cost plus contracts provide an incentive to shift cost and increase profits. Further, McGowan and Vondrzyk (2002) examine firms' government segment profitability relative to commercial segment profitability. Before 1998 and the implementation of Statement of Financial Accounting Standards 131, firms had more discretion in their segment reporting (Berger and Hann 2007) and could have had proprietary reasons to withhold reporting profitable segments. Thus, examining segment profitability also biases against finding differential profitability.

Second, we examine research and development (R&D) expense and selling, general, and administrative (SG&A) expense as the expenditures used by contractors to shift costs. Thomas and Tung (1992) examine defense contractors and whether they shift pension costs across contracts and time to the government. They find evidence of cost shifting using pension costs. One difficulty in interpreting their results is that cost plus contracts could provide an incentive to increase promised pension cost benefits. Thus, their observed cost increase might be due to increasing promised pension payments rather than cost shifting. In contrast, R&D and SG&A are not susceptible to this criticism. Our results confirm that procurement contractors shift costs, consistent with Thomas and Tung (1992), and provides evidence on other types of expenditures used for cost shifting purposes besides pension costs.

Third, we provide evidence that cost shifting is one explanation for the relatively higher profitability of government contractors. Lichtenberg (1992) finds excess profits on government contracts in the 1980s and suggests cost shifting as the explanation. McGowan and Vondrzyk (2002) do not examine cost shifting directly and infer from their segment profitability

comparisons that cost shifting is not a likely explanation for excess profitability. They suggest the result in Lichtenberg (1992) is likely due to other nonaccounting explanations. In contrast, we directly examine both cost shifting behavior and profitability. We find evidence that cost shifting is one explanation for relatively higher profitability.

Lastly, we are the first to provide evidence on the effectiveness of cost accounting standards. Our evidence suggests that cost accounting standards mitigate the positive association between cost plus contracts and profitability. When cost accounting standards apply to the contract, there is increased monitoring by the government and increased disclosure requirements about cost, which likely decrease the contractors ability to shift cost. Although our evidence sheds light on the benefits of cost accounting standards, we do not attempt to conclude on the tradeoff between the costs and benefits associated with these standards. The government must trade off the recovered gain from monitoring contractors with the cost of monitoring.

2. Background and hypotheses development

The contract terms under fixed price and cost plus contracts are different, which result in different incentives for the contractors. For a fixed price contract, the contractor provides a product or service to the government at a fixed price. For example, the department of homeland security contracts with Dell Inc. to buy laptop computers at a fixed price per unit. In most cases, no ex post renegotiation of the agreed upon price is permitted, therefore the government knows the final price before the project begins. Since the contract price is fixed, contract revenue is not sensitive to the seller's cost of production. The seller bears the risk associated with any cost overruns associated with the project.

With cost plus contracts the contract revenue is sensitive to the seller's cost. The contract revenue is equal to the seller's cost to produce the product or service plus a fixed fee or

guaranteed profit margin. Thus the government does not know the final price before the project begins and bears the risk of any cost overruns. Generally, when the costs are difficult to estimate ex ante or when the product or service is hard to explicitly define ahead of time the government may have incentives to offer a cost plus contract. For example, an aerospace firm contracts with the National Aeronautics and Space Administration (NASA) to build a rocket propulsion unit. For the aerospace company to produce the unit, they need to make a transaction specific investment in technology and human capital specifically tied to producing the rocket propulsion unit because the product is highly specialized and only in demand by NASA. From the aerospace firms perspective engaging in this transaction under a fixed price contract is perhaps too risky given the required amount of investment and the inherent uncertainty and complexity of the project. To appropriately share risk and incentivize the aerospace firm to make the requisite amount of investment, NASA would likely have to offer a cost plus contract and bear the risk of any cost overruns.

Under cost plus contracts, contractors have incentives to shift costs to the government whereas under fixed price contracts they generally do not. Obviously, if incurred costs were observable, there could be no cost shifting, the government would reimburse the contractor for expenditures incurred only for the contracted project. However, when there is information asymmetry between the contractor and the government, the true cost information is difficult to verify. Popular opinion often holds that cost plus contracts enable the contractors to cost shift and help the contractors earn a higher profitability. Since the contractor's revenue increases in the cost reimbursed, the contractor's profit will be higher when it shifts other cost to the government contracts. Specifically, if a firm receives cost plus contracts in some years and fixed

price contracts in others, its profitability could vary across years due to the differential availability of cost shifting based on contract type. This leads to our first hypothesis:

H1: Government procurement contractors' profitability is higher in years with a cost plus contract relative to years without cost plus contracts.

Thomas and Tung (1992) point out that contractor firms can shift costs both inter-contract and inter-time. With inter-contract shifting, firms shift cost from non-cost plus contracts (e.g., private sector or fixed price government contracts) to cost plus contracts within the same contract period. For example, a firm spends a total of \$100 million on R&D expenditures and reports \$100 million in R&D expense on the income statement. Out of \$100 million, \$20 million is incurred for the government contract and \$80 million was for the rest of the commercial business. However, the contractor firm could shift \$5 million of R&D expenses from commercial business to the government project. Therefore, the government must reimburse the firm \$25 million for its R&D expenses, while the firm's commercial section gets "subsidized" by \$5 million. Empirically, it is difficult to observe inter-contract cost shifting because in both cases the firm reports \$100 million in R&D expense on the income statement.

With inter-period shifting, firms shift costs into contract periods with a cost plus contract from periods without a cost plus contract. For example, a firm spends a total of \$100 million on R&D expenditures. Let us again assume the contractor firm should incur \$20 million of R&D expenses on the government project. With inter-period cost shifting the firm could purposefully engage in more R&D activities than necessary to fulfill the government contract, and these research activities could benefit its commercial business in the future. If the incremental expense amounts to an additional \$5 million in R&D costs, the government ends up subsidizing the contractor's commercial business by overpaying \$5 million dollars. With inter-period cost

shifting the firm reports \$105 million on the income statement for the year it has a cost plus contract with the government, instead of \$100 million if it does not. Therefore inter-period cost shifting can be detected empirically from the firm's financial statement information.

In this paper, we proxy for the government contractors' inter-period cost shifting using their discretionary expenditures. Our second hypothesis is therefore:

H2: Government procurement contractors' discretionary expenditure is higher in years with a cost plus contract relative to years with only fixed price contracts.

When there is information asymmetry between the contractor and the government, the contractor can include more expenditures in reimbursed cost than were actually incurred. Since the contractor's revenue is increasing in cost, more cost shifting will lead to a greater profitability. However, the government can increase its monitoring efforts to prevent cost shifting. With increased monitoring, the government can decrease the ability of contractors to shift costs. As such, we expect the association between cost contracts and profitability to decrease as monitoring by the government increases.

In terms of increased monitoring, some government contracts are subject to Cost Accounting Standards (CAS). Cost Accounting Standards are a set of 19 standards and rules for use in determining costs.⁴ A contract could be subject to full CAS coverage (i.e. required to follow all 19 standards), modified CAS coverage (required to follow only four standards: CAS #401, #402, #405, and #406)⁵, or be exempt from coverage. There are many reasons a contract could be exempt from CAS, including that a contract is less than \$7.5 million (and the firm does

⁴ For example, Thomas and Tung (1992) describe the CAS pension rules (#412 and #413) as similar to tax rules and Generally Accepted Accounting Principles (GAAP) but more strict than Employee Retirement Income Security Act (ERISA) rules.

⁵ CAS #401 "Consistency in Estimating, Accumulating and Reporting Costs"; CAS#402 "Consistency in Allocating Costs Incurred for the Same Purpose"; CAS #405 "Accounting for Unallowable Costs; CAS #406 "Cost Accounting Period".

not have other contracts over 7.5 million), a contract is for commercial items, a contract has “adequate price competition”, etc. When contracts are subject to full CAS coverage, the firm is required to provide a CAS Disclosure Statement, which describes in detail the accounting policies used to determine CAS related costs.

CAS standards can apply for both cost plus contracts and fixed price contracts. For fixed price contracts the information subject to CAS could be used in determining the contract price (i.e. the government uses the information to set the fixed price). For cost plus contracts the information subject to CAS could be used in determining the fixed fee, profit margin, or reimbursed costs. To the extent, contracts subject to CAS indicate increased monitoring efforts by the government to prevent cost shifting, we expect the association between cost contracts and profitability to decrease as monitoring by the government increases. This leads to our third hypothesis:

H3: Government contractors with cost plus contracts are relatively less profitable when cost accounting standards apply.

3. Research Design

3.1. Measure of discretionary expenditure Discretionary expenditures are defined as the sum of R&D, advertising, and SG&A expenditures. We estimate the following model based on the entire Compustat sample with data available to estimate the models following Roychowdhury (2006). We estimate the normal level of discretionary expenditures as follows:

$$DisExp_t / Assets_{t-1} = \mu_0 + \mu_1 (I / Assets_{t-1}) + \mu_2 (S_{t-1} / Assets_{t-1}) + \varepsilon_t \quad (1)$$

where,

$DisExp_t$ = research and development expense plus advertising and selling, general and administrative expense
 $Assets_{t-1}$ = total assets
 S_t = sales

Model (1) is estimated cross-sectionally for each industry-year with at least 15 observations. Continuous variables are winsorized at the 1% and 99% level to reduce the influence of outliers. For each firm-year, abnormal discretionary expenditure (*Abnormal DisExp*) is the residual from the corresponding industry-year regression. As an alternative proxy for abnormally high discretionary expenditures, we create an indicator variable equal to one if the residual from model (1) is in the highest quintile and zero otherwise (*Abnormal DisExp_Q10*) for a given year.

3.2. Hypothesis testing

To test our first hypothesis (H1), we adopt the following regression model used in Core et al. (1999):

$$ROA = \beta_0 + \beta_1 \text{Cost contract} + \beta_2 \text{Lagged ROA} + \beta_3 \text{StdROA} + \beta_4 \text{Log sales} + \beta_4 \text{Log procurement sales} + \varepsilon_t \quad (2)$$

where,

StdROA = standard deviation of ROA during the prior three years including year t

Log sales = the natural logarithm of sales in millions

Log procurement sales = the natural logarithm of procurement contract sales in millions

We include year and industry (two-digit SIC) fixed effects. We include the standard deviation of ROA over the fiscal year (*Std ROA*) and the natural logarithm of annual net sales (*LnSales*). We augment the Core et al. (1999) regression model and include *ROA* to control for the time-series properties of performance. Of particular interest to us is the coefficient on *Cost contract*. If contracting firms have higher return-on-assets in years with cost plus contracts, we should observe a positive β_1 .

To test our second hypothesis (H2), we estimate the following OLS regression model:

$$\text{Abnormal DisExp/Abnormal DisExp_Q10} = \alpha_0 + \alpha_1 \text{Cost contract} + \alpha_2 \text{Log assets} + \alpha_3 \text{Market-to-book} + \alpha_4 \text{ROA} + \varepsilon_t \quad (3)$$

where,

<i>Abnormal DisExp</i>	= the residual from model (1)
<i>Abnormal DisExp_Q10</i>	= an indicator variable equal to one if the residual from model (1) is in the highest quintile, and zero otherwise
<i>Cost contract</i>	= an indicator variable equal to 1 if the firm has at least one cost contract, and zero otherwise
<i>Log asset</i>	= natural logarithm of total assets in millions
<i>Market-to-book</i>	= the market value of equity divided by the book value
<i>ROA</i>	= earnings before extraordinary items divided by total assets

In all regressions, continuous variables are winsorized at the 1% and 99% level to reduce the influence of outliers and the standard errors are corrected to control for clustering across firm and year (Gow et al. 2010; Petersen 2009). The control variables are from Gunny (2010). Firm size, *Log assets*, controls for any size effects and the market-to-book ratio controls for growth opportunities. Return on assets, *ROA*, is included to address concerns that abnormal discretionary expenditures are correlated with performance. Of particular interest to us is the coefficient on *Cost contract*, α_1 . If contracting firms have higher discretionary expenditures in years with cost plus procurement contracts, we should observe a positive α_1 . A positive coefficient would be consistent with inter-period cost shifting and the results in Thomas and Tung (1992).

To test our third hypothesis (H3), we augment model (3) as follows:

$$ROA = \gamma_0 + \gamma_1 Cost\ contract_t + \gamma_2 \%CAS_t + \gamma_3 \%CAS * Cost\ contract_t + \gamma_4 Lagged\ ROA + \gamma_5 StdROA + \gamma_6 Log\ sales + \gamma_7 Log\ procurement\ sales + \varepsilon_t \quad (4)$$

where,

$\%CAS$ = percentage of procurement revenue subject to the cost accounting standards

Hypothesis three suggests that the positive association between whether the firm has a cost plus contract and performance will be constrained as the government's monitoring effort increases. Our proxy for monitoring is the percent of contract revenue that is subject to the cost accounting standards ($\%CAS$). When CAS standards apply, not only specific rules must be applied when determining costs but also the firm must produce CAS Disclosure Statement,

which describes in detail the accounting policies used to determine CAS related costs. CAS standards can apply to fixed and cost plus contracts. For fixed contracts, this could help reduce information asymmetry when negotiating the price of the contract. For cost plus contracts, CAS could help by preventing cost shifting. Therefore, we expect a negative coefficient on the variable *%CAS*. Of particular interest to us is the coefficient on the interaction between *Cost contract* and *%CAS*. If firms with cost plus contracts are less profitable when a greater percentage of contract revenue is subject to CAS, we would expect a negative coefficient on the interaction term, γ_3 .

4. Sample and Descriptive Statistics

4.1. Sample selection

Panel A of Table 1 describes the sample details. First, we retrieve data on all procurement contracts between 2004 and 2010 from the Federal Procurement Data System website. Next, we map the 21,657,724 contracts to firms available in Compustat and retain 6,098,655 contracts. Next, after requiring nonmissing control variables, we are left with 4,912,232 contracts, or 1,137 firms. Table 1 reveals that 209,278 (4.26%) of contracts are cost plus and 379 (33.33%) firms have at least one cost plus contract over the sample period. Panel B of Table 1 describes our three samples. Sample 1 consists of all procurement contractors in our sample including 5,745 firm-years (1,137 firms), of which 1,315 firm-years involve cost plus contracts and 4,439 firm-years do not. The next two samples exclude firms without cost plus contracts over the sample period. As discussed earlier, the government only offers cost plus contracts in certain situations, such as when the costs are difficult to estimate *ex ante* or when the product or service is hard to explicitly define ahead of time. As such, firms that receive cost contracts could be different than firms that do not receive cost contracts and this self-selection could bias our results.

Sample 2 excludes firms without any cost contracts over the sample period and consists of 2,201 firm-years (383 firms), of which 1,315 firm-years have cost plus contracts and 886 do not. Sample 3 consists of firms with at least one year with a cost plus contract and one year without over the sample period. This sample consists of 1,505 firm-years (258 firms), of which 619 firm-years have cost plus contracts and 886 do not.⁶ Sample 3 allows each firm to serve as its own control and mitigate concerns that self-selection is driving our findings. Therefore, we use sample 3 in our main analysis and conduct sensitivity analysis using sample 1 and sample 2.

4.2. Descriptive statistics

Panel A of Table 2 provides the descriptive statistics of contract revenue and frequency by contracting department. The first four columns describe all contracts and the last four columns provide analysis for cost plus contracts. The top three departments by contract revenue is the department of the army, navy, and air force.⁷ The general services administration and the department of veteran affairs have the greatest frequency of contracts.

Panel B of Table 2 provides the descriptive statistics of contract revenue and frequency by 3-digit North American Industry Classification System (NAICS). The government procurement database provides the NAICS code associated with each contract. The top three industries for all contracts by contract revenue are professional, scientific, and technical services (541), transportation equipment manufacturing (336), and computer and electronic product manufacturing (334). For cost plus contracts, professional, scientific, and technical services (541) is the largest and consists of 63% of all cost plus contract revenue. The next largest industry is transportation equipment manufacturing (336) which consists of 14% of cost plus

⁶ The difference between sample 2 and sample 3 is that firms with a cost contract every year ($383 - 258 = 125$ firms) are excluded.

⁷ Because of data availability prior studies typically use a sample consisting of the Top 100 department of defense contracts (by revenue) who are publicly traded (e.g., Thomas and Tung 1992). The department of defense consists of the department of the army, navy, air force, and defense (non military).

contract revenue. Cost plus contracts are typically awarded when the product or service is hard to explicitly define ahead of time. Consistent with this notion, industries with products and services that are difficult to specify (e.g., professional, scientific, and technical services) are awarded more frequently with cost plus contracts whereas industries with less complicated offerings (e.g., merchant wholesalers, nondurable goods) are not.

Table 3 presents descriptive statistics for a sample of 258 firms with at least one year with a cost plus contract and one year with only fixed price contracts. Panel A provides summary statistics for sample firm-years with at least one cost plus contract, and Panel B for firm-years with only fixed price contracts. Firm-years with cost contracts have significantly higher mean and median abnormal discretionary expenditures (*Abnormal DisExp*) than firm-years with only fixed price contracts. Also, both the mean and median return-on-assets (*ROA*) are significantly higher for firm-years with at least one cost plus contract compared to firm-years with only fixed contracts. Overall, the univariate differences indicate that firm-years with at least one cost contract have and better financial performance and more discretionary expenditures, consistent with hypothesis one and two.

Between the samples with cost plus contracts and fixed price contracts, mean assets, market-to-book, standard deviation of ROA, sales, and procurement sales are not significantly different. These insignificant differences between the two samples indicate that our sample selection procedure that uses the firm as its own control mitigates concerns that firm characteristics drive the variation in our variables of interest (i.e. abnormal discretionary expenditures and return-on-assets). Lastly, %CAS is significantly higher for the cost plus sample, which is expected since cost plus contracts are more likely to be subject to the cost accounting standards.

5. Results

5.1. Main results

We report the results of estimating model (2) in Table 4. If cost plus contracts are associated with higher profitability relative to fixed contracts, we would expect the contractors to have better financial performance in years with cost plus contracts. Consistent with this expectation, we find a significantly positive relation between return-on-assets and whether the firm has a cost contract. Specifically, the coefficient on *Cost contract*, is 0.008 ($p = 0.03$). This suggests that firm-years with cost plus contracts have higher profitability compared to firm-years with only fixed price contracts, consistent with our hypothesis one. In terms of economic significance, the marginal effect on *Cost contract* is 0.762%. The marginal effect can be interpreted as the percentage increase in *ROA* during years with a cost contract. Mean assets for the sample is \$3.33 million, therefore, firm-years with cost contracts have earnings before extraordinary items that is \$25.37 million higher than firm years without cost contracts, holding all other variables at their mean.

We report the results of estimating model (3) in Table 5. We use two variables to proxy for abnormal discretionary expenditures. First, we estimate an ordinary least squares regression with the continuous variable, *Abnormal DisExp*, as the dependent variable. Second, we estimate a logistic regression with the dichotomous variable, *Abnormal DisExp_Q10*, as the dependent variable. The first column reveals a significant relation between abnormal discretionary expenditures and whether the firm has a cost contract. Specifically, the coefficient on *Cost contract*, is 0.128 ($p = 0.04$). The second column reveals a significant relation between whether the firm is in the highest quintile of abnormal discretionary expenditures and whether the firm has a cost contract. The coefficient on *Cost contract*, is 0.411 ($p = 0.01$). The analysis presented

in Table 5 provides evidence consistent with hypothesis two, that procurement contractors shift costs across time to the government through cost plus contracts. Overall, our results provide evidence of cost shifting.

Taken together, the results in Table 4 and 5 suggest government procurement contractors shift costs to the government when they have cost plus contracts and this leads to greater profitability relative to firm-years with only fixed price contracts. Please note two important issues when interpreting these results. First, hypothesis two tests for inter-period cost shifting since it is easier to observe than inter-contract shifting. Even though we test exclusively for inter-period shifting, we believe both types of cost shifting could occur. Second, just because firms shift cost and have higher profits, we do not necessarily believe this is suboptimal from the governments perspective. The optimality of contract form is determined by many factors. Just because cost plus contracts may lead to distorted incentives, the appropriateness of fixed price contracts is still not warranted.

Hypothesis three suggests that the positive association between having a cost plus contract and the firm's financial performance is constrained as the government's monitoring increases. Our proxy for monitoring is the percent of contract revenue that is subject to cost accounting standards (*%CAS*). We test our third hypothesis (H3) and report the results of model (4) in Table 6. We continue to observe a positive relation between having a cost contract and the firm's return on assets. The coefficient on *%CAS* is negative and significant, suggesting firm profitability decreases as the percentage of contract revenue subject to cost accounting standards increases. More importantly, we show that the relation between *Cost contract* and *ROA* varies with *%CAS* in the predicted manner. Specifically, the coefficient on the interaction between *Cost contract* and *%CAS* is significantly negative ($\gamma_3 = -0.047$, $p = 0.08$). These results provide

support for our third hypothesis (H3) and suggest that as the government's monitoring increases the association between having a cost plus contract and profitability decreases.

5.2. Sensitivity Analysis

For our main analysis, our sample consists of the 258 firms that have at least one year with a cost plus contract and one year without, over the sample period (see Panel B of Table 1, sample 3). For this sample, the firm serves as its own control. For a sensitivity test, we explore the robustness of our results to alternative control samples. We estimate model (2) and model (3) using two alternative samples: all contractors (see Panel B of Table 1, sample 1) and all contractors with at least once cost contract over the sample period (see Panel B of Table 1, sample 2).

We report the result of sensitivity analysis of estimating model (2) in Panel A of Table 9. For both alternative control samples, we continue to find a positive association between abnormal discretionary expenditures and the firm having a cost plus contract. Specifically, the coefficient on *Cost contract* for Sample 1, is 0.0659 ($p = 0.05$) and the coefficient on *Cost contract* for Sample 2, is 0.0906 ($p=0.04$). Next, we report the result of sensitivity analysis of estimating model (3) in Panel B of Table 9. Again, we find a significantly positive relation between the firms' return on assets and the firms having a cost contract. Specifically, the coefficient on *Cost contract* for Sample 1, is 0.0045 ($p = 0.02$) and coefficient on *Cost contract* for Sample 2, is 0.0070 ($p = 0.03$). This provides evidence that our findings that firm-years with cost plus contracts have higher profitability compared to firm-years with only fixed price contracts is robust to alternative control samples.

We also use alternative categorization of contracts to re-run our tests. For the main analyses, we refer to any contracts with revenue that is sensitive to the seller's cost as cost plus.

However, Rogerson (1992) describes four common types of contracts employed by the U.S. Department of Defense: (1) pure fixed price (2) pure cost reimbursement (3) incentive fixed price (i.e. realized costs are reimbursed up to a pre-specified threshold level) (4) incentive cost reimbursement (i.e. incentive fixed price contracts revert to pure cost reimbursement at a pre-specified threshold cost level). We estimate model (2) and model (3) on the sample including only (1) pure fixed price and (2) pure cost reimbursement and the results are robust.

6. Conclusion

In government procurement, cost plus contracts facilitate risk-sharing between the government and the contractors, especially when the product or service is difficult to define ex-ante. It is however also believed that contractors abuse the flexibility provided by cost plus contracts and shift other costs to the government to earn a higher profit than deserved. In this paper, we use a unique dataset of the U.S. federal procurement contracts to examine whether contractors that are awarded with cost plus contracts earn a higher profit, and whether the higher profit is associated with cost shifting behavior. We find evidence consistent with cost plus contract being associated with higher profitability and greater discretionary expenditures, which are indicative of cost shifting. We also find that the government's monitoring through the Cost Accounting Standards helps mitigate this problem.

Our findings however do not indicate that cost plus contracts are inferior to fixed price contracts. Optimal contract form is determined by many factors, such as project-specific risk, market competition, contractor characteristics, and etc. Cost plus contracts are necessary when the uncertainty associated with the project is high. Further, once the cost plus contract is chosen, the contractors can earn an informational rent since the government cannot observe the true cost

incurred by the contractor. More stringent monitoring is very costly and can be more expensive to the tax payers.

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Table 1

Panel A: Sample selection

Year	All federal procurement contracts	Contracts mapped to Compustat	Contracts with control variables	Firms	Cost plus contracts	Firms
2005	2,849,310	685,199	495,573	861	32,998	237
2006	3,723,011	926,604	753,137	916	35,445	247
2007	4,008,419	1,167,812	972,510	937	33,079	264
2008	4,391,358	1,335,355	1,114,443	1,000	31,242	237
2009	3,341,336	1,019,336	799,357	1,016	34,704	231
2010	3,344,290	964,349	777,212	1,024	41,810	249
	21,657,724	6,098,655	4,912,232	1,137	209,278	379

Panel B: Control samples

	Firm-years	Firms	Firm-years with at least one cost contract	Firm-years with no cost contract (i.e. only fixed)
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Sample 1 5,745 1,137 1,315 4,439

*Includes all procurement contractors

Sample 2 2,201 383 1,315 886

*Includes procurement contractors with at least one cost contract over the sample period

Sample 3 1,505 258 619 886

*Includes procurement contractors with at least one year with a cost plus contract and one year with only fixed contracts

Table 2

Panel A: Contract revenue and frequency by contracting department

Department	All contracts				Cost plus contracts			
	Contract revenue (in millions)	%	# of contracts	%	Contract revenue (in millions)	%	# of contracts	%
Department of the army	75,820	25%	318,283	6%	40,855	36%	99,010	47%
Department of the navy	68,641	23%	234,272	5%	20,072	18%	24,277	12%
Department of the air force	62,234	21%	144,796	3%	27,682	24%	45,030	22%
General services administration	15,311	5%	1,333,167	27%	1,118	1%	1,047	1%
Department of veteran affairs	13,858	5%	1,135,401	23%	45	0%	183	0%
National Aeronautics and space admin.	10,725	4%	27,088	1%	7,587	7%	8,471	4%
Department of homeland security	9,565	3%	16,292	0%	3,999	3%	4,426	2%
Department of defense (non military)	8,046	3%	46,589	1%	2,154	2%	1,979	1%
Department of health and human services	7,388	2%	42,961	1%	2,895	3%	3,264	2%
Department of the treasury	4,934	2%	24,265	0%	1,799	2%	2,463	1%
Department of justice	3,606	1%	78,920	2%	131	0%	156	0%
Department of the interior	2,925	1%	54,899	1%	371	0%	672	0%
Department of state	2,530	1%	8,684	0%	824	1%	1,249	1%
Department of transportation	2,421	1%	23,819	0%	167	0%	173	0%
Department of energy	2,382	1%	7,209	0%	1,647	1%	1,475	1%
Department of commerce	1,978	1%	18,546	0%	665	1%	1,467	1%
Railroad retirement board	1,368	0%	4,455	0%	1	0%	7	0%
Department of agriculture	1,327	0%	13,020	0%	935	1%	5,858	3%
Agency for international development	1,011	0%	1,825	0%	682	1%	462	0%
Department of the labor	861	0%	9,501	0%	422	0%	442	0%
Other	3,427	1%	1,368,240		408		7,167	
	300,358		4,912,232		114,458		209,278	

Panel B: Contract revenue and frequency by industry

NAICS Industry	All contracts				Cost plus contracts			
	Contract revenue (in millions)	%	# of contracts	%	Contract revenue (in millions)	%	# of contracts	%
Professional, scientific, and technical services	108,586	36%	252,848	5%	71,808	63%	133,448	64%
Transportation equipment manufacturing	57,522	19%	429,885	9%	15,884	14%	21,356	10%
Computer and electronic product manufacturing	35,906	12%	234,442	5%	6,858	6%	8,215	4%
Merchant wholesalers, nondurable goods	17,785	6%	454,306	9%	10	0%	57	0%
Administrative and support services	10,272	3%	32,511	1%	5,680	5%	5,402	3%
Merchant wholesalers, durable goods	7,246	2%	345,302	7%	54	0%	82	0%
Telecommunications	7,202	2%	64,842	1%	779	1%	1,099	1%
Fabricated metal product manufacturing	6,588	2%	133,014	3%	575	1%	691	0%
Machinery manufacturing	6,353	2%	87,216	2%	744	1%	1,229	1%
Air transportation	5,119	2%	8,566	0%	1	0%	29	0%
Chemical manufacturing	4,036	1%	120,157	2%	96	0%	79	0%
Miscellaneous manufacturing	3,859	1%	167,507	3%	61	0%	155	0%
Repair and maintenance	3,825	1%	25,850	1%	1,790	2%	2,483	1%
Electronics and appliance stores	3,205	1%	44,088	1%	130	0%	201	0%
Food manufacturing	2,897	1%	10,249	0%	2	0%	4	0%
Waste management and remediation service	2,770	1%	15,920	0%	1,308	1%	1,758	1%
Educational services	2,765	1%	14,246	0%	645	1%	576	0%
Utilities	2,569	1%	10,032	0%	32	0%	113	0%
Electrical equipment, appliance, and component manufact.	2,129	1%	29,662	1%	165	0%	260	0%
Furniture and related product manufacturing	2,101	1%	46,645	1%	6	0%	140	0%
Other	7,624	3%	2,384,944		7,831		31,901	
	300,358		4,912,232		114,458		209,278	

Table 3

Descriptive statistics by contract type

Panel A: Firm-years with cost plus pricing contracts for sample of firms with one cost-plus year and one fixed control only year (619 firm-years and 258 firms)

	Mean	Median	Std. Dev.	25%	75%
Abnormal DisExp	0.36	0.00	1.73	-0.22	0.33
ROA	0.04	0.05	0.13	0.02	0.09
Assets	8.08	8.09	2.10	6.78	9.62
Market-to-book	2.94	2.14	28.19	1.38	3.87
Std ROA	0.03	0.02	0.06	0.01	0.04
Sales	7.93	8.02	1.97	6.78	9.37
Procurement sales (in millions)	65.96	14.49	174.01	3.32	47.83
%CAS	0.05	0.00	0.16	0.00	0.00

Panel B: Firm-years with only fixed price contracts for sample of firms with one cost-plus year and one fixed price contract only year (886 firm-years and 258 firms)

	Mean	Median	Std. Dev.	25%	75%
Abnormal DisExp	0.18**	-0.02**	1.36	-0.20	0.31
ROA	0.02**	0.04***	0.14	0.01	0.08
Assets	8.13	8.16	2.20	6.66	9.82
Market-to-book	2.69	2.03*	9.65	1.33	3.45
Std ROA	0.04	0.02	0.06	0.01	0.03
Sales	7.93	8.00	2.05	6.75	9.47
Procurement sales (in millions)	57.08	6.29*	177.19	1.35	30.52
%CAS	0.02***	0.00***	0.12	0.00	0.00

The sample consists of 1,505 firm-year observations between 2005 and 2010. Panel A includes 619 firm-years (258) firms with at least one cost plus contract. Panel B includes 886 firm-years (258 firms) with only fixed price contracts. ***/**/* represent statistical significance at 1%/5%/10% levels between the means and medians of the sample in Panel A compared to the sample in Panel B. All continuous variables are winsorized at the 1% and 99% level. *Abnormal DisExp* is the residual from model (1). *Log assets* is the natural logarithm of total assets in millions. *Market-to-book* is the market value of equity divided by the book value. *ROA* is earnings before extraordinary items divided by total assets. *StdROA* is the standard deviation of ROA during the prior three years including year *t*. *Log sales* is the natural logarithm of sales in millions. *Procurement sales* is procurement contract sales in millions. *%CAS* is the percentage of procurement revenue subject to the cost accounting standards.

Table 4

Regression results of return on assets on whether the firm has cost plus contracts

	Pred.	ROA _t	p-value
Intercept		-0.030	(0.35)
Cost contract	+	0.008	(0.03)
Lagged ROA	+	0.595	(0.00)
Std ROA	-	-0.372	(0.27)
Log sales	+	0.008	(0.05)
Log procurement sales	+	-0.001	(0.32)
Year and industry indicators		Yes	
No. of observations		1,505	
R ²		0.606	

The sample consists of 1,505 firm-year observations between 2005 and 2010. The regression includes year and industry (two-digit SIC) indicator variables. The standard errors are clustered by firm and year. Two-tailed *p* values are reported in parentheses. ***/**/* represent statistical significance at 1%/5%/10% levels. All continuous variables are winsorized at the 1% and 99% level. *ROA* is earnings before extraordinary items divided by total assets. *Cost contract* is an indicator variable equal to 1 if the firm has at least one cost contract, and zero otherwise. *StdROA* is the standard deviation of ROA during the prior three years including year *t*. *Log sales* is the natural logarithm of sales. *Log procurement sales* is the natural logarithm of procurement contract sales.

Table 5

Regression results of abnormal discretionary expense on whether the firm has cost plus contracts and controls

	Abnormal DisExp	p-value	Abnormal DisExp_Q10	p-value
Intercept	0.230	(0.24)	-2.991	(0.00)
Cost contract	0.128	(0.04)	0.411	(0.01)
Log assets	-0.029	(0.22)	-0.071	(0.22)
Market-to-book	0.000	(0.54)	-0.002	(0.38)
ROA	0.268	(0.06)	-0.123	(0.74)
Year and industry indicators	Yes		Yes	
No. of observations	1,505		1,505	
R ²	0.023			
Pseudo R ²			0.114	

The sample consists of 1,505 firm-year observations between 2005 and 2010. The regression includes year and industry (two-digit SIC) indicator variables. The standard errors are clustered by firm and year. Two-tailed *p* values are reported in parentheses. ***/**/* represent statistical significance at 1%/5%/10% levels. All continuous variables are winsorized at the 1% and 99% level. *Abnormal DisExp* is the residual from model (1). *Abnormal DisExp_Q10* is an indicator variable equal to one if the residual from model (1) is in the highest quintile, and zero otherwise. *Cost contract* is an indicator variable equal to 1 if the firm has at least one cost contract, and zero otherwise. *Log asset* is the natural logarithm of total assets in millions. *Market-to-book* is the market value of equity divided by the book value. *ROA* is earnings before extraordinary items divided by total assets. The regression with *Abnormal DisExp* is estimated using ordinary least squares. The regression with *Abnormal DisExp_10Q* is estimated using a logistic regression.

Table 6

Regression results of return on assets on whether the firm has cost plus contracts and subject to cost accounting standards

	Pred.	ROA _t	p-value
Intercept		-0.025	(0.39)
Cost contract	+	0.012	(0.02)
%CAS	-	-0.059	(0.05)
Cost contract * %CAS	-	-0.047	(0.08)
Lagged ROA	+	0.579	(0.00)
Std ROA	-	-0.373	(0.26)
Log sales	+	0.008	(0.05)
Log procurement sales	+	-0.001	(0.42)
Year and industry indicators		Yes	
No. of observations		1,505	
R ²		0.613	

The sample consists of 1,505 firm-year observations between 2005 and 2010. The regression includes year and industry (two-digit SIC) indicator variables. The standard errors are clustered by firm and year. Two-tailed *p* values are reported in parentheses. ***/**/* represent statistical significance at 1%/5%/10% levels. All continuous variables are winsorized at the 1% and 99% level. *ROA* is earnings before extraordinary items divided by total assets. *Cost contract* is an indicator variable equal to 1 if the firm has at least one cost contract, and zero otherwise. *StdROA* is the standard deviation of ROA during the prior three years including year *t*. *Log sales* is the natural logarithm of sales. *Log procurement sales* is the natural logarithm of procurement contract sales. *%CAS* is the percentage of procurement revenue subject to the cost accounting standards.

Table 7

Sensitivity analysis for alternative control samples

Panel A: Regression results of abnormal discretionary expense on whether the firm has cost plus contracts and controls

	Pred.	Abnormal DisExp	p-value	Highest quintile of abnormal DisExp	p-value
Intercept		0.4660	(0.01)	0.4205	(0.19)
Cost contracts	+	0.0659	(0.05)	0.0906	(0.04)
Log assets	+	-0.0506	(0.00)	-0.0339	(0.07)
Market-to-book	-	0.0004	(0.33)	0.0001	(0.85)
ROA	+	0.2938	(0.03)	0.3139	(0.04)
No. of observations		5,754		2,201	
R ²		0.088		0.061	

Panel B: Regression results of return on assets on whether the firm has cost plus contracts and cost accounting standards clause

		All contractors		All contractors with at least one cost contract	
	Pred.	ROA _t	p-value	ROA _t	p-value
Intercept		0.1245	(0.20)	0.0044	(0.82)
Cost contract	+	0.0045	(0.02)	0.0070	(0.03)
Lagged ROA	+	0.5511	(0.00)	0.5396	(0.00)
Std ROA	-	-0.2087	(0.08)	-0.4501	(0.17)
Log sales	+	0.0093	(0.00)	0.0061	(0.03)
Log procurement sales	+	-0.0004	(0.55)	-0.0019	(0.13)
Year and industry indicators		Yes		Yes	
No. of observations		5,754		2,201	
No. of firms		1,137		383	
R ²		0.44		0.57	

The sample consists of 1,505 firm-year observations between 2005 and 2010. The regressions include year and industry (two-digit SIC) indicator variables. The standard errors are clustered by firm and year. Two-tailed p values are reported in parentheses. ***/**/* represent statistical significance at 1%/5%/10% levels. All continuous variables are winsorized at the 1% and 99% level. *Abnormal DisExp* is the residual from model (1). *Abnormal DisExp_Q10* is an indicator variable equal to one if the residual from model (1) is in the highest quintile, and zero otherwise. *Cost contract* is an indicator variable equal to 1 if the firm has at least one cost contract, and zero otherwise. *Log asset* is the natural logarithm of total assets in millions. *Market-to-book* is the market value of equity divided by the book value. *ROA* is earnings before extraordinary items divided by total assets. *Cost contract* is an indicator variable equal to 1 if the firm has at least one cost contract, and zero otherwise. *StdROA* is the standard deviation of ROA during the prior three years including year t . *Log sales* is the natural logarithm of sales. *Log*